

## Knowledge is Power? Outcome Probability Information Impairs Detection of Deceptive Intent

### Highlights

- We examined the effect of outcome probability information on detection of deception
- Probability information more strongly biased responses to deceptive actions
- Probability information more strongly biased responses of low-skilled players
- Knowledge of player tendencies impaired detection of deceptive intent

## **Abstract**

The benefits and costs of prior expectations that are (i.e., congruent) or are not in harmony (i.e., incongruent) with action outcomes appear to be balanced; however, researchers have yet to examine the influence on skilled detection of deception. In this study we investigated whether response bias resulting from probability information (a) is stronger for low-skilled than high-skilled participants, (b) is stronger for deceptive actions than genuine actions, and (c) impairs the discriminability of genuine and deceptive actions. High-skilled ( $n = 15$ ) and low-skilled ( $n = 15$ ) soccer players responded to life-sized projected video clips showing an oncoming opponent taking the ball to their left or right, with or without a deceptive ‘stepover’ action. Three probability conditions were used with respect to outcome direction: 50/50, 67/33, and 83/17. Participants responded by stepping on one of two corresponding pressure mats, as if attempting to intercept the player. Response accuracy for genuine and deceptive actions was used to generate measures of bias ( $c$ ) and sensitivity ( $d'$ ). The results confirmed stronger probability bias for deceptive actions than genuine ones, and for low-skilled than high-skilled participants. Congruence between high outcome probability and the direction of the fake significantly enhanced the effectiveness of the deceptive action. The study provides the first evidence that outcome probability information impairs skilled detection of deceptive intent.

*Keywords:* deception; anticipation; soccer

# **Knowledge is Power? Outcome Probability Information Impairs Detection of Deceptive Intent**

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## Author contribution statement

Robin Jackson: Conceptualization, Methodology, Formal analysis, Data Curation, Writing – Original draft, Supervision. Hayley Barton: Investigation, Resources, Data Curation, Writing – Review & Editing, Project administration. Daniel Bishop: Resources, Writing – Review & Editing, Supervision

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9 **Abstract**

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46 *Keywords:* deception; anticipation; soccer  
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51 **Introduction**  
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62 Expert performers are highly attuned to opponent kinematics, which enables them to  
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64 anticipate action outcomes better than their less-skilled counterparts (Mann, Williams, Ward,  
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66 & Janelle, 2007; Müller & Abernethy, 2012). Researchers have shown that experts also use a  
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68 variety of static and dynamic sources of contextual information to form accurate expectations  
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70 about the most likely action outcome (Williams & Jackson, 2019). These include the position  
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72 of players relative to each other, the ball, and the court in racket sports (Abernethy, Gill,  
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74 Parks, & Packer, 2001; Murphy, Jackson, Cooke, Roca, Benguigui, & Williams, 2016), the  
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76 sequence of preceding actions and action outcomes (Farrow & Reid, 2012; Gray, 2002;  
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78 Loffing, Stern, & Hagemann; 2015; Murphy, Jackson, & Williams, 2019a), game situation  
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80 and fielder positions in cricket (Runswick, Roca, Williams, McRobert, & North, 2019), and  
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82 individual player tendencies or action preferences (Gredin, Bishop, Broadbent, Tucker, &  
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84 Williams, 2018; Mann, Schaeffers, & Cañal-Bruland, 2014; Milazzo, Farrow, Ruffault, &  
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86 Fournier, 2016). However, few researchers have addressed the question of how contextual  
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88 and kinematic sources of information combine to affect anticipation (Cañal-Bruland & Mann,  
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90 2015; Loffing & Cañal-Bruland, 2017). Moreover, while one study showed that head fakes  
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92 were less effective when they were used with high relative frequency (Güldenpenning,  
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94 Alaboud, Kunde, & Weigelt, 2018) researchers have not examined how action outcome  
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96 expectations affect skilled detection of deceptive intent (Güldenpenning, Kunde, & Weigelt,  
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98 2017; Murphy, Jackson, & Williams, 2019b). This is a crucial question because a body of  
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100 research shows that the ability to detect deception is highly diagnostic of expertise,  
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102 characterised by earlier differentiation of genuine and deceptive actions and better  
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104 suppression of responses to deceptive signals (Brault, Bideau, Kulpa, & Craig, 2012; Jackson,  
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106 Warren, & Abernethy, 2006; Jackson, Barton, Ashford, & Abernethy, 2018).  
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121 Accordingly, the purpose of the present study is to examine how action outcome probability  
122 information affects anticipation and detection of deception. We employed signal detection  
123 analysis to measure the extent to which outcome probability information biased participants'  
124 responses and affected their ability to discriminate between genuine and deceptive actions.  
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126 Alongside its theoretical contribution, the practical significance of this study is underscored  
127 by the pervasiveness of deceptive actions in interactive duelling sports coupled with the  
128 widespread use of performance analysis to identify player preferences. For example,  
129 approximately half of all successful receptions in international netball matches were preceded  
130 by a deceptive action (Fox, Spittle, Otago, & Saunders, 2014) and 72% of successful tackle-  
131 breaks in rugby union were preceded by a sidestep (Wheeler, Askew, & Sayers, 2010).  
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#### 145 Situational probability and anticipation

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149 The benefits of prior expectations that are congruent (i.e., in harmony) with action outcomes  
150 are well established. In a seminal study, Alain and Proteau (1977) recorded response times as  
151 participants struck one of two tennis balls suspended next to a light on their left and right.  
152 Before each block of trials, participants were told the probability that each light would  
153 illuminate, which ranged from 1.0/0 to 0.1/0.9. The researchers found that on congruent trials  
154 (i.e., when the higher-probability light illuminated) response times were faster for extreme  
155 values (0.9/0.1; 0.1/0.9; 1.0/0) than in the equal probability condition (0.5/0.5). Similar  
156 benefits of congruent high probability information were shown in studies that indirectly  
157 manipulated participant expectations. Skilled baseball batters made faster decisions,  
158 committed fewer errors, and timed their swings more accurately when contextual 'pitch  
159 count' information (e.g., balls, strikes, outs, runners on base) allowed them to identify the  
160 more likely pitch (Gray, 2002; Paull & Glencross, 1997). Researchers have also shown that  
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180 higher-skilled performers are better at detecting statistical regularities. Tennis players  
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182 progressively responded more quickly and accurately to the first serve of each block of eight  
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184 when it was consistently directed to the same place, and karate players did so when an attack  
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186 sequence was repeated every fourth trial (Farrow & Reid, 2012; Milazzo, Farrow, Ruffault, &  
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188 Fournier, 2016).  
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193 A key question regarding congruence concerns the balance between the benefits and costs of  
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195 congruent and incongruent action outcomes, respectively. Handball goalkeepers improved  
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197 their ability to anticipate penalty throw direction but only when the thrower showed the same  
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199 action preference as in training (Mann, Schaefers, & Cañal-Bruland, 2014). Conversely,  
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201 anticipation accuracy *decreased* when the thrower showed a different action preference in the  
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203 post-test than in training. If the costs and benefits are balanced there is a mathematical  
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205 advantage in knowing situational probabilities because, by definition, congruent action  
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207 outcomes are more frequent. Evidence from studies of genuine actions supports this  
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209 proposition. Relative to a ‘no pre-cue’ control condition, baseball batter response times were  
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211 faster and response accuracy was slightly higher following a valid pre-cue, while response  
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213 times were slightly slower and response accuracy was lower following an invalid pre-cue  
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215 (Radlo, Janelle, Barba, & Frehlich, 2001). Likewise, a simulated two-on-two soccer task  
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217 revealed that high-skilled players with knowledge of the action preference of the player in  
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219 possession of the ball to dribble (67%) or pass to their teammate (33%) were more accurate  
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221 on congruent trials and less accurate on incongruent trials, relative to when they had no prior  
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223 information (Gredin et al., 2018). A field-based study of soccer goalkeeper responses to  
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225 penalty kicks similarly found equivalent costs and benefits for incongruent and congruent  
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227 action outcomes (Navia, van der Kamp, & Ruiz, 2013). On half of the trials, the ball was  
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229 directed to the goalkeeper’s left and right with equal probability. For the remaining trials there  
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239 was a 0.8/0.2 probability of the ball being kicked to each side of the goal. The goalkeepers  
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241 dived to the correct side on approximately half of kicks across the equal probability and no  
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243 probability conditions. The unequal probability condition led to an equivalent increase (high-  
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245 probability side) and decrease (low-probability side) in the proportion of saves and dives to  
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247 the correct side of the goal. Other studies in which there was no clear control condition  
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249 nonetheless showed clear differences between congruent and incongruent action outcomes. In  
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251 baseball, bat swing timing error for fast pitches was significantly greater when the pitch  
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253 followed three slow pitches (incongruent sequence) than when it followed three fast pitches  
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255 (congruent sequence, Gray, 2002). Similarly, skilled volleyball players responded more  
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257 quickly and accurately when the outcome was congruent with the four preceding trials but  
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259 were slower and less accurate when it was incongruent (Loffing et al., 2015).  
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265 Drawing on measures from signal detection theory (Swets, Tanner, & Birdsall, 1961),  
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267 researchers have argued that prior expectations will lead to cognitive strategies that bias  
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269 responses on visual discrimination tasks but will not affect discriminability (Cañal-Bruland,  
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271 Filius, & Oudejans, 2015). In support, the researchers showed that baseball batters tended to  
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273 ‘sit on a fastball’; in other words, to bat as if expecting a fastball. Extending this account to  
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275 perceptual judgment tasks that require different responses to left and right action outcomes,  
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277 probability information would be expected to bias responses to genuine and deceptive actions  
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279 equally, but have no effect on their discriminability. An alternative proposal, based on  
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281 confirmation bias, is that prior expectations lead performers to place greater weight on  
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283 kinematic information that confirms their expectation (Runswick et al., 2019). These  
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285 researchers reasoned that skilled players would be more susceptible to performance  
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287 decrements associated with incongruent action outcomes because they made greater use of  
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289 contextual information. In support, skilled cricket batsmen were better than less-skilled  
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298 batsmen at judging the outcome of deliveries (the point where the ball passed the stumps) but  
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300 only when it was congruent with fielder positions, bowler type and game context. Conversely,  
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302 high-skilled players performed worse than less-skilled players when outcome was incongruent  
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304 with context.  
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309 Deceptive actions provide an opportunity to test these alternative accounts because attempts  
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311 to ‘sell the dummy’ are often characterised by *exaggeration* in order to elicit an incorrect  
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313 response. For example, foot displacement, head yaw and upper trunk yaw were found to be  
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315 exaggerated in sidestep actions compared to actions without sidesteps (Brault, Bideau, Craig,  
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317 & Kulpa, 2010). In soccer stepovers, evidence of exaggeration was seen in participant  
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319 responses to early-occluded video clips, which resulted in more incorrect responses than the  
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321 equivalent correct responses to genuine actions. This led to a stronger reverse spatial cueing  
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323 effect than the positive effect observed for genuine actions (Jackson et al., 2018; Wright &  
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325 Jackson, 2014). Spatial exaggeration increases the distinctiveness of images and actions,  
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327 which in turn makes them easier to identify; for example, caricatures of famous faces tend to  
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329 be recognised faster than veridical portraits (Mauro & Kubovy, 1992). In regard to action  
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331 recognition, researchers found that different styles of tennis serve were more accurately  
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333 identified when their defining kinematic characteristics were exaggerated (Pollick, Fidopiastis  
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335 & Braden, 2001). Similarly, soccer goalkeepers anticipated the direction of non-deceptive  
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337 penalty kicks more accurately when the kicker ‘telegraphed’ their intentions by exaggerating  
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339 their kicking action (Smeeton & Williams, 2012). Consistent with this property, we  
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341 hypothesise that situational probability information will more strongly affect responses to  
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343 deceptive actions than genuine actions, thereby increasing susceptibility to deception when  
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345 the fake movement is aligned with outcome expectations. If correct, strong expectations will  
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347 result in a greater increase in ‘false alarm’ responses to deceptive actions than correct  
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357 responses to genuine actions. This, in turn, will impair the discriminability of genuine and  
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359 deceptive actions.  
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364 To test this prediction, we used signal detection measures derived from the proportion of  
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366 correct responses to one type of stimulus (i.e., genuine actions) and the corresponding  
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368 proportion of ‘false alarm’ responses to the other type of stimulus (i.e., deceptive actions) in  
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370 order to measure the discriminability of genuine and deceptive actions ( $d'$ ) and participant  
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372 bias in favour of judging an action to be genuine or deceptive ( $c$ ). These indices take into  
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374 account the possibility that participants might have different levels of bias toward judging an  
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376 action to be deceptive. To illustrate with an extreme example, an individual might achieve  
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378 100% accuracy when judging deceptive actions. However, until we know their accuracy on  
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380 genuine trials we don't know to what extent this score represents perfect discrimination  
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382 between genuine and deceptive actions (they score 100% on both: high discriminability, zero  
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384 bias) or an extreme bias toward judging every action to be deceptive (they score 0% on  
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386 genuine trials: high bias, zero discriminability). In other domains, signal detection analysis  
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388 has been used to examine deception in facial expressions and personal injury narratives  
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390 (Porter, Juodis, ten Brinke, Klein, & Wilson, 2010), judgments of truth and deception in law  
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392 enforcement (Meissner & Kassin, 2002), and in comparisons of methods of eye-witness  
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394 identification (Wixted & Mickes, 2014). In sport research, signal detection analysis revealed  
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396 that skilled handball goalkeepers and outfield players were equally good at differentiating  
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398 genuine and deceptive penalty throws but that the goalkeepers showed a stronger bias toward  
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400 judging throws to be deceptive (Cañal-Bruland & Schmidt, 2009). Conversely, high-skilled  
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402 soccer players were much better than low-skilled players at differentiating genuine and  
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404 deceptive actions but both groups showed a bias to judge actions to be genuine (Jackson et al.,  
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406 2018; Wright, Bishop, Jackson, & Abernethy, 2013).  
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418 In the present study we sought to test four specific hypotheses regarding the effect of outcome  
419 probability information on dynamic judgments of deceptive intent. Our first hypothesis (H1)  
420 is that high-skilled soccer players will be better than low-skilled players at discriminating  
421 between genuine and deceptive actions. Our second hypothesis (H2) is that probability  
422 information will more strongly bias responses of low-skilled performers than high-skilled  
423 performers, due to their relative inability to differentiate genuine and deceptive actions. In  
424 line with evidence for stronger responses to exaggerated actions, our third hypothesis (H3) is  
425 that probability information will more strongly bias responses to deceptive actions than  
426 genuine actions. As a result of this imbalance, our fourth hypothesis (H4) is that probability  
427 information will impair the discriminability of genuine and deceptive actions relative to the  
428 equal probability condition.  
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444 To test the first two hypotheses, we defined ‘hits’ as correct responses to genuine actions and  
445 ‘false alarms’ as incorrect responses to deceptive actions (Jackson et al., 2018). To test H3  
446 and H4, we defined ‘hits’ as correct responses when the player took the ball to the high  
447 probability (congruent) side and ‘false alarms’ as incorrect responses when the player took the  
448 ball to the low-probability (incongruent) side. This second analysis must be conducted  
449 separately for genuine and deceptive actions and in this case response bias measures the  
450 tendency to respond in line with the higher probability value (negative  $c$  values) or lower  
451 probability value (positive  $c$  values). These values are then compared against the equal  
452 probability condition in which, by definition, there is no bias. The index of discriminability  
453 ( $d'$ ) measures the ability to discriminate between left and right action outcomes in each  
454 probability condition.  
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## Method

**Participants.** A medium effect size ( $f = 0.25$ ) with Power set at 0.8 for two groups and five levels of the repeated measure yielded a recommended total sample size of 22 for the within-factors and within-between interaction, and 28 for the tests involving three levels of the repeated measure. A total of 30 male soccer players were recruited. High-skilled participants ( $n = 15$ ) were semi-professional players with a mean age of 20.6 years ( $SD = 2.4$ ) and a mean of 14.2 years ( $SD = 2.8$ ) of competitive playing experience. The low-skilled participants ( $n = 15$ ) had a mean age of 24.4 years ( $SD = 4.1$  years) and a mean of 1.3 years ( $SD = 1.2$ ) of competitive playing experience at recreational level (e.g., local Sunday league club). Institutional ethical approval was granted and all participants gave written consent prior to participating in the study.

**Test construction.** We used a video camera (Canon HV40, Tokyo, Japan, 50 frames per second) mounted on a tripod at a height of 1.5 m to record high-definition sequences of two male right-footed players (hereafter ‘actors’) from an English Premier League academy. Each actor ran with the ball directly towards the video camera from a distance of 11.5 m. They then changed direction to the left or right as they would when trying to evade another player. For non-deceptive actions, the actors simply manoeuvred the ball to the left or right of the camera with their lead foot. For the deceptive actions, the actors executed a ‘stepover’ by moving their lead foot in front of and across the ball before taking it in the opposite direction. The actors were instructed to perform as they would in a real one-on-one situation, running at pace directly towards the camera while maintaining close control of the ball. A panel of two UEFA B-license and one UEFA A-license coaches rated each video clip for the player’s smoothness of movement, pace of the approach, and execution of the direction change. Overall mean

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534 scores were used to identify the 12 best examples from each player, which were used to  
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536 develop the test stimuli.  
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541 The test comprised 12 blocks of 12 trials, containing six blocks from each actor. Actor order  
542 was counterbalanced across participants. The test trials for each actor comprised two blocks  
543 of 12 trials in each of the three levels of situational probability information (50/50, 67/33, and  
544 83/17); these values indicated the percentage likelihood that the actor would change direction  
545 to the left or right, respectively. For the 67/33 and 83/17 conditions, the higher probability  
546 value was associated with leftward direction changes for one block and rightward changes for  
547 the other. The order of presentation of the three levels of probability was counterbalanced  
548 across participants and was identical for the two actors. Each trial was occluded 80 ms after  
549 the lead foot contacted (genuine trials) or would have contacted (deceptive trials) the ball and  
550 was followed by a 5-second inter-trial interval. The test stimuli were presented using E-Prime  
551 software and were projected (Optoma HD25, CA, USA) to create a 1.6 m (h) x 2.1 m (w)  
552 scene. Participants stood 2.8 m from the screen and responded by stepping on one of two  
553 response mats (0.72 x 0.39 m, Defender Security, Farnell, Leeds, UK) placed on the floor in  
554 front, and to the left and right, of the participant.  
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572 **Procedure.** Participants were told they would view brief video clips designed to replicate a  
573 one-on-one scenario in which a player in possession runs towards them then takes the ball to  
574 their left or right. They were told that the actor in the video would simply take the ball to their  
575 left or right on half of the trials, and on the remaining trials would fake taking the ball one  
576 way before going the other by executing a 'stepover'. Participants were told that their task  
577 was to judge whether the actor would take the ball to their left or right and that they should  
578 respond by stepping on the left or right response pad. They were instructed to respond as  
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591 quickly and accurately as possible as they would in an actual one-on-one situation.

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595 Participants were informed that they would view 12 blocks of 12 trials and that the chance of  
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597 the actor taking the ball to the left or right would be indicated on the screen before each block  
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599 and would be either 50/50, 67/33 (or 33/67), or 83/17 (or 17/83). To familiarise participants  
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601 with the experiment setup they were shown 16 trials, comprising four genuine and four  
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603 deceptive trials from each actor. These video clips were different from those used in the test  
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605 and probability information was not provided. During the test, the chance of the actor taking  
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607 the ball to the left and right was displayed on the left and right side of the screen for 13  
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609 seconds before each block of 12 trials. Participants were given a 20-second break after each  
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611 block of trials and were given a two-minute break between the blocks of trials for the first  
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613 actor (Blocks 1-6) and the second actor (Blocks 7-12).  
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618 **Data analysis.** Our first analysis focused on the extent to which different probability values  
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620 (1) affected the ability to discriminate between genuine and deceptive actions, and (2) biased  
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622 participants to judge actions to be genuine or deceptive. For this analysis, the index of  
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624 discriminability ( $d'$ ) was calculated by subtracting  $z$ -scores for the proportion of incorrect  
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626 responses on deceptive trials ('false alarms') from  $z$ -scores for the proportion of correct  
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628 responses on genuine trials ('hits'). More positive values indicate better discrimination  
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630 between genuine and deceptive actions. The criterion for responding ( $c$ ) was calculated by  
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632 taking the average of the  $z$ -scores for 'hits' and 'false alarms'. Negative values indicate a bias  
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634 to judge actions to be genuine and positive values indicated a bias to judge actions to be  
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636 deceptive. The  $d'$  and  $c$  indices were calculated for each of the five outcome probability  
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638 values associated with the initial directional intention (left or right) conveyed by the action  
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640 (0.17, 0.33, 0.50, 0.67, 0.83).  
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652 The second analysis focused on (1) the effect of each probability pair (50/50, 67/33, and  
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654 83/17) on the discriminability of left and right action outcomes, and (2) the extent to which  
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656 participant responses were biased in favour of the higher (or lower) probability value. For this  
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658 analysis, the index of discriminability ( $d'$ ) was calculated by subtracting  $z$ -scores for the  
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660 proportion of incorrect responses on low probability outcomes ('false alarms') from  $z$ -scores  
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662 for the proportion of correct responses on high probability outcomes ('hits'). The criterion for  
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664 responding ( $c$ ) was calculated by taking the average of the  $z$ -scores for 'hits' and 'false  
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666 alarms'.  
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672 To prevent infinite  $z$ -scores that result from scores of 0 and 1, these values were replaced with  
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674 0.021 and 0.979, respectively, using the formulae  $1/2n$  and  $(n - 0.5)/n$  (Stanislaw & Todorov,  
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676 1999), where  $n$  represents the number of normal and deceptive trials in the three probability  
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678 conditions. The maximum value of  $d'$  was  $\pm 4.07$  and for  $c$  was  $\pm 2.03$ . Alpha was set at .05 for  
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680 all analyses and partial eta squared ( $\eta_p^2$ ) was used to indicate effect size. The Greenhouse-  
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682 Geisser adjustment to the degrees of freedom was applied when Mauchly's test of sphericity  
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684 was violated.  
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689 Last, we recorded response time to examine a potential trade-off between response time and  
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691 response accuracy. For each analysis the assumptions relating to parametric analyses and the  
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693  $F$  distribution, such as normality, homogeneity of variances, independence of raw scores, and  
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695 sphericity of the repeated measures values were evaluated. The univariate output was assessed  
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697 with alpha set at .05 and the Greenhouse-Geisser correction was applied to the degrees of  
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699 freedom when the sphericity assumption was violated.  
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## Results

The proportion of correct responses to genuine and deceptive actions for each of the five outcome probability values is shown in Figure 1, which has four features of interest. First, the high-skilled group (HS; square symbols) was much more accurate overall than the low-skilled group (LS; triangle symbols). Second, the difference in response accuracy for genuine trials (solid lines) and deceptive trials (dotted lines) was much greater in the LS group than the HS group. Third, as the outcome probability value associated with the initial direction change increased so too did the proportion of correct responses to genuine trials, whereas the proportion of correct responses to deceptive trials decreased. Fourth, the effect of the probability information appeared to be stronger for the LS group than the HS group.

**\*\*Figure 1 about here\*\***

### **Effect of outcome probability on the discriminability of genuine and deceptive actions**

The response sensitivity data across the five action outcome probabilities are shown in Figure 2. The HS group was substantially better than the LS group at discriminating between genuine and deceptive actions; that is, these participants made more ‘hits’ (correct responses on genuine trials) coupled with fewer ‘false alarm’ responses (incorrect responses on deceptive trials). Analysis of variance confirmed a significant and large effect of expertise,  $F(1, 28) = 54.76, p < .001, \eta_p^2 = .66$ , and a significant effect of outcome probability,  $F(2.79, 78.19) = 6.25, p = .001, \eta_p^2 = .18$ . High probability values resulted in a greater increase in ‘false alarm’ responses on deceptive trials than ‘hits’ on genuine trials (Figure 1). As a result, there was a notable decrease in the discriminability of genuine and deceptive actions for high probability



768  
 769  
 770 values relative to the 50% condition. This was reflected in the results of simple contrast tests,  
 771  
 772 which showed that sensitivity was significantly impaired in the 83% condition,  $F(1, 28) =$   
 773  
 774  $10.32, p = .003$ , and not in the 17% ( $p = .95$ ), 33% ( $p = .13$ ), or 67% ( $p = .10$ ) conditions.  
 775  
 776 High-skilled players had a consistent advantage over less-skilled players across all five  
 777  
 778 probability values, which was borne out by a non-significant Expertise x Probability  
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 780 interaction,  $F(2.79, 78.19) = 0.40, p = .74, \eta_p^2 = .01$ .  
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 785 The response bias data ( $c$ ) for the HS and LS groups are shown in Figure 3. We found strong  
 786  
 787 support for the hypothesised relationship between outcome probability value and the degree  
 788  
 789 of bias toward judging an action to be genuine,  $F(2.39, 66.85) = 30.46, p < .001, \eta_p^2 = .52$ .  
 790  
 791 Moreover, the significant and large effect of expertise showed that the probability information  
 792  
 793 had a stronger effect on the the LS group than the HS group,  $F(1, 28) = 20.66, p < .001, \eta_p^2 =$   
 794  
 795  $.43$ . The different slopes yielded a significant Expertise x Probability interaction,  $F(2.39,$   
 796  
 797  $66.85) = 3.22, p = .04, \eta_p^2 = .10$ . One-tailed independent  $t$ -tests showed that bias toward  
 798  
 799 judging actions to be genuine was significantly stronger in the LS group than the HS group  
 800  
 801 for the 33% ( $p = .04, 95\% \text{ CI} = -0.02 \text{ to } 0.44$ ), 50% ( $p = .001, 95\% \text{ CI} = 0.11 \text{ to } 0.42$ ), 67% ( $p$   
 802  
 803  $< .001, 95\% \text{ CI} = 0.26 \text{ to } 0.77$ ), and 83% ( $p = .002, 95\% \text{ CI} = 0.19 \text{ to } 0.95$ ) probability values.  
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809 \*\*Figure 2 about here\*\*

810 \*\*Figure 3 about here\*\*

### 811 812 813 814 **Bias toward responding to the high-probability side**

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 817 The degree to which probability information biased participant responses to genuine and  
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 819 deceptive actions is shown in Figure 4, in which values of zero for the 50/50 condition are  
 820  
 821 included for reference. It is clear that the probability information biased participants to  
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829 respond to the high-probability side and this bias increased from the 67/33 to the 83/17  
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831 condition. This was reflected in a significant effect of probability condition,  $F(1, 28) = 9.91, p$   
832  
833  $= .004, \eta_p^2 = .26$ . Response bias was clearly stronger for deceptive actions than genuine  
834  
835 actions in both groups and this was reflected in a significant effect of deception,  $F(1, 28) =$   
836  
837  $9.94, p = .004, \eta_p^2 = .26$ . Consistent with analysis of the individual outcome probability  
838  
839 values, bias to respond to the high-probability side was stronger in the LS group than the HS  
840  
841 group,  $F(1, 28) = 5.06, p = .03, \eta_p^2 = .15$ . All other effects were non-significant.  
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846 The effect of probability information on task performance in genuine and deceptive trials is  
847  
848 illustrated in Figure 5. Consistent with the data shown in Figure 1, the HS group was  
849  
850 substantially better than the LS group at discriminating between left and right action  
851  
852 outcomes,  $F(1, 28) = 53.88, p < .001, \eta_p^2 = .66$ . The significant and large effect of deception  
853  
854 confirmed that participants were much better at judging genuine actions than deceptive ones,  
855  
856  $F(1, 28) = 87.13, p < .001, \eta_p^2 = .66$ . There was also a significant Expertise x Deception  
857  
858 interaction,  $F(1, 28) = 21.03, p < .001, \eta_p^2 = .43$ , which was caused by a stronger effect of  
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860 expertise on judgments of deceptive actions (HS:  $M = 3.23$ ; LS:  $M = 0.77$ ; difference = 2.46)  
861  
862 than genuine actions (HS:  $M = 3.85$ ; LS:  $M = 2.58$ ; difference = 1.27). Last, there was a  
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864 significant effect of probability,  $F(2, 56) = 4.41, p = .02, \eta_p^2 = .14$ , which was caused by  
865  
866 significantly lower discriminability scores in the 83-17 condition ( $M = 2.41, SE = 0.13$ ) than  
867  
868 in the 50-50 condition ( $M = 2.72, SE = 0.16$ ) and the 67-33 condition ( $M = 2.69, SE = 0.15$ ).  
869  
870 Of particular note, poorer discriminability for deceptive actions in the 83-17 condition  
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872 reflected a larger increase in ‘false alarm’ responses when fakes were directed to the high-  
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874 probability side than the equivalent decrease in ‘false alarm’ responses for fakes to the low-  
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876 probability side (see Figure 1).  
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888 In sum, the signal detection analyses clearly show that probability information biased  
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890 participant responses, that the effect was stronger for low-skilled than high-skilled  
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892 performers, and that it was stronger for deceptive actions than genuine actions. As a result, the  
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894 ability to discriminate between genuine and deceptive actions declined as probability values  
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896 increased.  
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899 \*\*Figure 4 about here\*\*  
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902 \*\*Figure 5 about here\*\*  
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### 904 905 **Response time**

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907 Analysis of response times for genuine and deceptive trials across the five outcome  
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909 probability values revealed that high-skilled participants ( $M = 1734$  ms,  $SE = 26$  ms)  
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911 responded faster than low-skilled participants ( $M = 1810$  ms,  $SE = 26$  ms),  $F(1, 28) = 4.32$ ,  $p$   
912  
913  $= .047$ ,  $\eta_p^2 = .13$ . This indicates there was no speed-accuracy trade-off: the high-skilled group  
914  
915 was both faster and more accurate. The analysis also revealed a small but statistically  
916  
917 significant Expertise x Probability interaction,  $F(4, 112) = 2.48$ ,  $p = .048$ ,  $\eta_p^2 = .08$ . As can be  
918  
919 seen in Figure 6, response times for the HS group were consistent across probability values  
920  
921 whereas response times of the LS group were slightly slower for the lower probability values.  
922  
923 There was no difference in response times for genuine actions ( $M = 1773$  ms,  $SE = 18$  ms)  
924  
925 and deceptive actions ( $M = 1771$  ms,  $SE = 20$  ms),  $F(1, 28) = 0.06$ ,  $p = .81$ ,  $\eta_p^2 = .00$ .  
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932 \*\*Figure 6 about here\*\*  
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### 934 935 **Discussion**

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947 In recent review papers researchers highlighted the rudimentary understanding of how  
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949 situational probability information affects anticipation and perception of deceptive intent  
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951 (Jackson & Cañal-Bruland, 2019; Müller & Abernethy, 2012; Murphy et al., 2019b). The  
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953 purpose of the present study was to begin to address this by determining how outcome  
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955 probability information affects judgments of deceptive intent. To do so we analysed the  
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957 degree to which probability information biased high-skilled and less-skilled participants'  
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959 responses, and its effect on their ability to differentiate genuine and deceptive actions. The  
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961 evidence in support of our central hypotheses was compelling. First, high-skilled participants  
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963 were better than low-skilled counterparts at discriminating between genuine and deceptive  
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965 actions (Jackson et al., 2018; Wright & Jackson, 2014). In line with our second hypothesis,  
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967 response bias associated with situational probability information was stronger for low-skilled  
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969 participants than high-skilled participants (see Figure 3). In both groups, probability  
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971 information biased participant responses in a consistent manner, characterised by a linear  
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973 relationship between the response criterion ( $c$ ) and the outcome probability value (Figure 3).  
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975 In simple terms, when the actor shaped to take the ball to the left or right, the tendency of  
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977 participants to judge the action to be genuine changed as a function of outcome probability  
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979 information. In line with our third hypothesis, the same probability information biased  
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981 responses to deceptive actions more strongly than genuine ones and this was true for both  
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983 high-skilled and low-skilled participants (see Figure 4). In line with our fourth hypothesis,  
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985 this impaired discriminability of genuine and deceptive actions relative to the neutral (50/50)  
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987 condition (Figure 5).  
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994 In regard to expertise, researchers have shown that experts make more effective use of  
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996 dynamic contextual information than less-skilled players (Gredin, et al., 2018). In the present  
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998 study we sought to examine how different levels of fixed prior probability information  
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1006 affected high- and low-skilled performers' judgments of deceptive intent. The results clearly  
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1008 show that performance on the primary task influenced the degree of bias afforded by prior  
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1010 probability information. Low-skilled participants were worse than high-skilled participants at  
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1012 discriminating between genuine and deceptive actions; correspondingly, low-skilled  
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1014 participants were more strongly influenced by prior probability information. This finding is  
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1016 consistent with a recent study that showed situational probability information more strongly  
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1018 influenced judgments when kinematic information was ambiguous (Helm, Cañal-Bruland,  
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1020 Mann, Troje, & Munzert, 2020). In this study, the researchers systematically manipulated  
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1022 handball throw kinematics by applying a morphing procedure to create seven levels of  
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1024 kinematic ambiguity from exaggerated genuine actions (thrower released the ball) to  
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1026 exaggerated disguised actions (thrower retained the ball). Their results showed a curvilinear  
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1028 relationship between the degree of kinematic ambiguity and the effect of action preference  
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1030 information, with the effect strongest when the kinematic information was a 50-50 morph of  
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1032 genuine and disguised actions.  
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1038 The present study replicates elements of these findings and extends them in several respects.  
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1040 First, we manipulated situational probability in regard to directional outcome (left, right)  
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1042 rather than preference for using genuine and disguised actions, which was held constant at  
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1044 50%. Accordingly, our results extend the findings from action *preference* to action *outcome*  
1045  
1046 probabilities, which are arguably more typical of the information provided in many direction  
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1048 judgment tasks in sport (Navia et al., 2013). Second, we used high-skilled and low-skilled  
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1050 performers rather than novices. This allowed us to compare the effects of action outcome  
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1052 probability information on performers who differed in their ability to distinguish between  
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1054 genuine and deceptive actions. Importantly, the evidence from comparisons between the two  
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1056 groups corroborates and extends the findings of Helm et al. (2020), in that it supports the  
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1065 integration of kinematic and situational probability information in high-skilled as well as low-  
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1067 skilled performers, and in that the effect of probability information was stronger in performers  
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1069 who were less able to differentiate genuine and deceptive actions. Third, we occluded videos  
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1071 after the critical time-point (i.e., after the point at which the foot contacted or would have  
1072  
1073 contacted the ball) and instructed participants to make physical responses similar to those they  
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1075 would make when defending against a real player. This revealed perhaps the most striking  
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1077 finding, which was that high-skilled performers did extremely well on deceptive trials when  
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1079 the probability information was balanced (50/50 condition) but were nonetheless significantly  
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1081 impaired when strong expectations aligned with the (false) intention conveyed by a deceptive  
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1083 action (see Figure 1). Researchers have shown that a key characteristic of expert judgments of  
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1085 deceptive intent is their ability to suppress incorrect responses to deceptive actions (Brault et  
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1087 al., 2012). We found evidence of this in the 50/50 condition, yet the ability was significantly  
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1089 and progressively impaired as outcome probabilities associated with the direction of the fake  
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1091 strengthened. In a physical sense, the probability information changed the criterion for  
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1093 responding, resulting in fewer suppressed incorrect responses. To quantify this effect in  
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1095 practical terms, interpolation of the data in Figure 4 suggests that the response bias for  
1096  
1097 genuine actions in the 83/17 probability condition could be attained in deceptive actions with  
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1099 probabilities of just 61/39 (high-skilled group) and 65/35 (low-skilled group).  
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1106 It is important to note that the potential dual upsides of increased accuracy on high probability  
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1108 genuine trials and low probability deceptive trials were smaller in the high-skilled group than  
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1110 the low-skilled group because their baseline performance (50/50 condition) was higher. To  
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1112 confirm that probability information more strongly biased responses for low-skilled  
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1114 performers and for deceptive actions we inspected the relative *downsides* of more ‘false  
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1116 alarm’ responses on high probability deceptive trials and fewer ‘hits’ on low probability  
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1124 genuine trials, neither of which were subject to potential ceiling effects (Figure 1). In support  
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1126 of the stronger effect of probability information on low-skilled performers, response accuracy  
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1128 relative to the 50/50 condition decreased more in the low-skilled group than the high-skilled  
1129  
1130 group in both instances. Likewise, in support of a stronger effect of probability information on  
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1132 deceptive actions, the decrease in response accuracy on high probability deceptive trials was  
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1134 greater than the decrease in response accuracy on low probability genuine trials.  
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1139 In regard to the question of whether prior probability information leads to a cognitive bias or  
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1141 affects the way individuals interact with the task, the present results clearly show that  
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1143 probability information changed the judgment criterion more strongly for deceptive actions  
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1145 than genuine actions. As a result, participants incurred greater costs by way of more ‘false  
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1147 alarm’ responses to deceptive actions. These outweighed the benefits of more correct  
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1149 responses to genuine actions, which impaired their ability to discriminate between genuine  
1150  
1151 and deceptive actions. This provides further evidence that prior probability information and  
1152  
1153 kinematic information are integrated as the action unfolds (Gray & Cañal-Bruland, 2018;  
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1155 Helm, et al., 2020). More importantly, the way they did so was consistent with confirmation  
1156  
1157 bias, because response bias was stronger for the exaggerated confirmatory signals associated  
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1159 with deceptive actions than for the non-exaggerated signals in genuine actions (Runswick et  
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1161 al., 2019). We add a caveat to this interpretation, in that we did not analyse the kinematics of  
1162  
1163 the video stimuli used in the test. Exaggeration has been identified in other deceptive actions  
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1165 such as rugby sidesteps and deceptive soccer penalty kicks, and is considered integral to  
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1167 ‘selling’ the dummy. Nonetheless, systematic manipulation of the degree of exaggeration  
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1169 used to convey deceptive intent is needed to demonstrate unequivocally its role in  
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1171 strengthening biasing effects. From a signal detection perspective, we speculate there will be  
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1173 a point at which the exaggerated action becomes so different to the equivalent genuine action  
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1183 that action outcome can be readily identified so will be less effective in inducing confirmation  
1184 bias. Establishing this threshold, and how it is affected by prior expectations, will further  
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1186 enhance our understanding of how kinematic and contextual sources of information are  
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1189 integrated.  
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### 1191 1192 1193 Practical Implications 1194 1195 1196 1197

1198 The present study has two key practical implications. First, we question the efficacy of  
1199 providing ‘player preference’ information for action outcomes in skills in which deceptive  
1200 actions are common. Our results show that the performance benefits when action outcomes  
1201 were congruent with expectations were outweighed by the costs incurred when the performer  
1202 faked to move in the expected direction. In considering whether or not to provide such  
1203 information, it is instructive that the key finding – namely, deceptive actions became super-  
1204 deceptive when intent was aligned with high expectations – held true for both high-skilled and  
1205 low-skilled participants. High-skilled participants responded correctly to almost 95% of  
1206 deceptive trials in the neutral (50/50) condition yet this decreased to 76% when the fake intent  
1207 was aligned with strong outcome expectations (Figure 1, 83% probability). The potential  
1208 benefits of prior probability information were clearly greater for low-skilled performers  
1209 because they performed at a lower level (65%) in the 50/50 condition; however, this was  
1210 offset by their greater susceptibility to deception. Indeed low-skilled participants responded  
1211 correctly on just 35% of trials when deceptive intent was aligned with the strongest  
1212 expectations (Figure 1).  
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1232 Relatedly, a second practical implication is that performers should consider how they might  
1233 exploit the expectations of their opponent. Researchers have shown that unexpected action  
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1242 outcomes can be highly successful (Runswick, et al., 2019). The present study shows that by  
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1244 aligning deceptive intent conveyed by their actions with the outcome expectations of their  
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1246 opponent a performer can dramatically increase the potency of their deceptive action. This  
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1248 ‘higher order’ level of conditional deception can be used to exploit situations in which context  
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1250 leads to strong expectations regarding likely actions. It suggests, for example, that when  
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1252 player and court or field positions lead to a strong expectation on the part of a defensive  
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1254 player, the attacking player will be more successful by first faking to do the expected rather  
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1256 than simply doing the unexpected.  
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## 1261 Conclusion

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1263 The present findings offer compelling evidence for the influence of prior expectations on the  
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1265 perception of deceptive intent and lead to three main conclusions. First, the influence of  
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1267 situational probability information was evident in both high-skilled and low-skilled  
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1269 performers, which highlights the potent effect of prior expectations, particularly when they  
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1271 align with the intention conveyed by a fake action. Second, the stronger influence of  
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1273 situational probability information on deceptive actions than genuine actions supports its  
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1275 active integration during the task in a manner that is consistent with confirmation bias. Third,  
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1277 the stronger influence of situational probability information on low-skilled than high-skilled  
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1279 performers corroborates findings regarding the different weighting given to this information  
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1281 depending on the difficulty of the primary task. In sum, the results lead us to caution that  
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1283 situational probability information might harm anticipation in certain sports skills because  
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1285 deceptive actions become ‘super-deceptive’ when aligned with observer expectations.  
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1307  
1308 **Competing interests**  
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1310 There are no competing interests to declare.  
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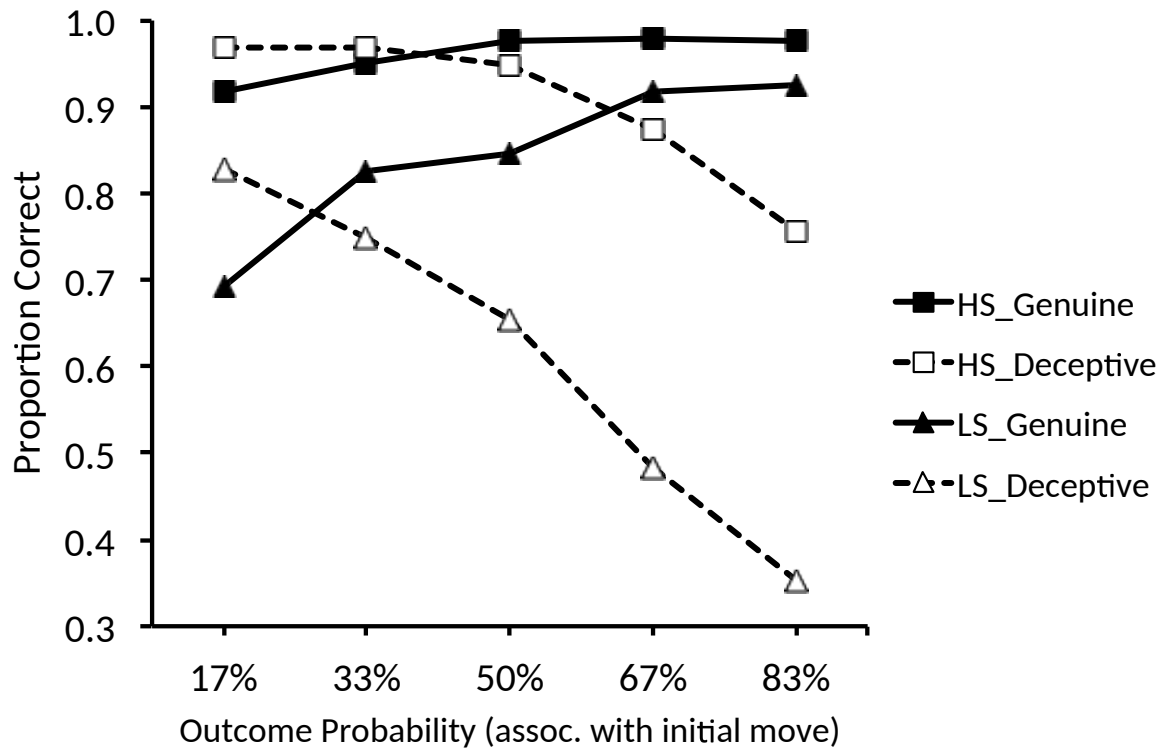
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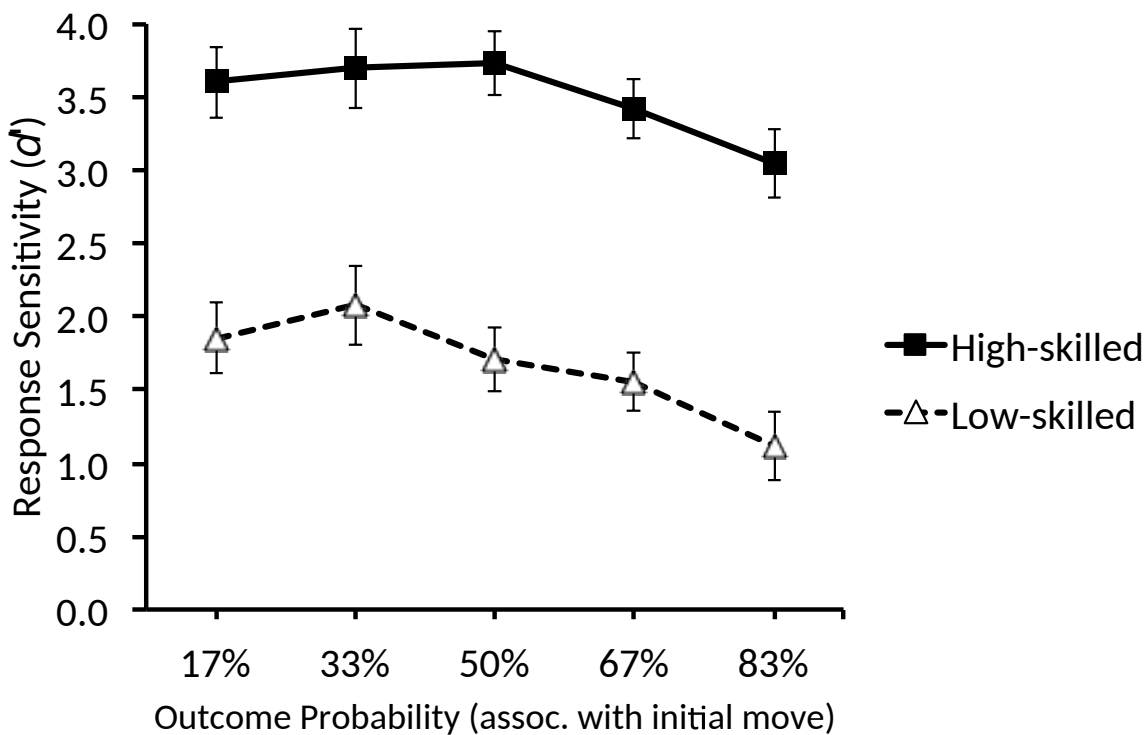
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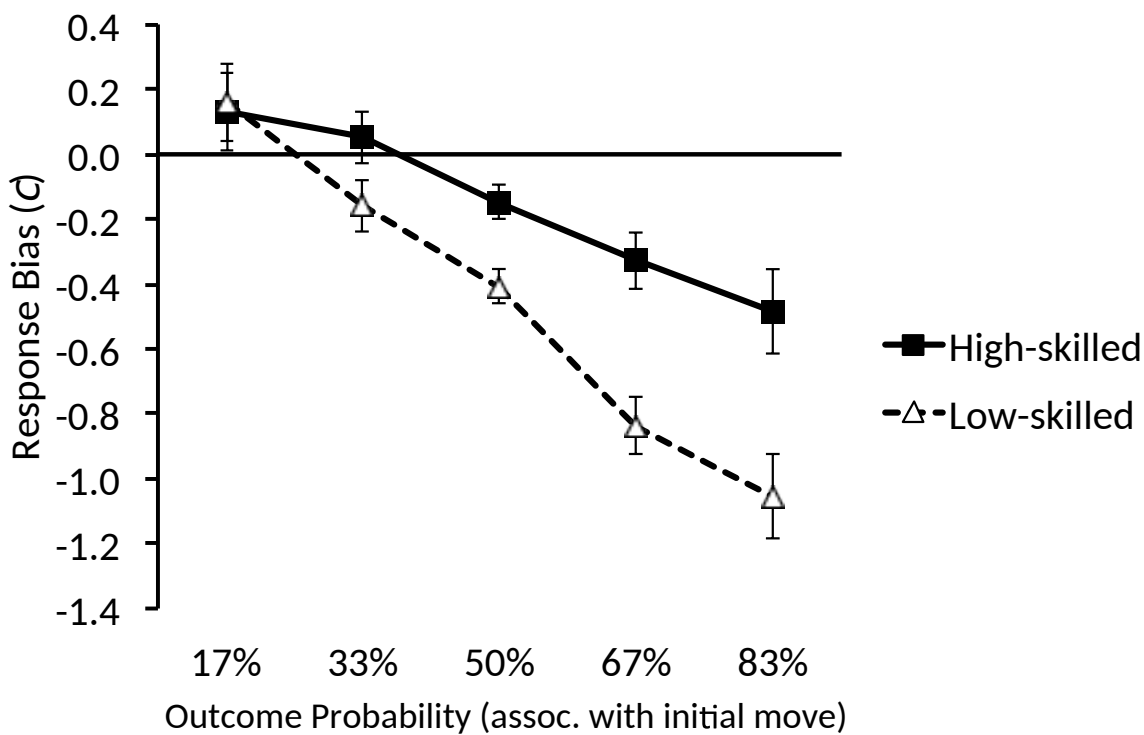


**Figure 1.** The proportion of correct responses made by the high-skilled (HS) and low-skilled (LS) group on genuine (solid lines) and deceptive (dashed lines) trials for each outcome probability value.

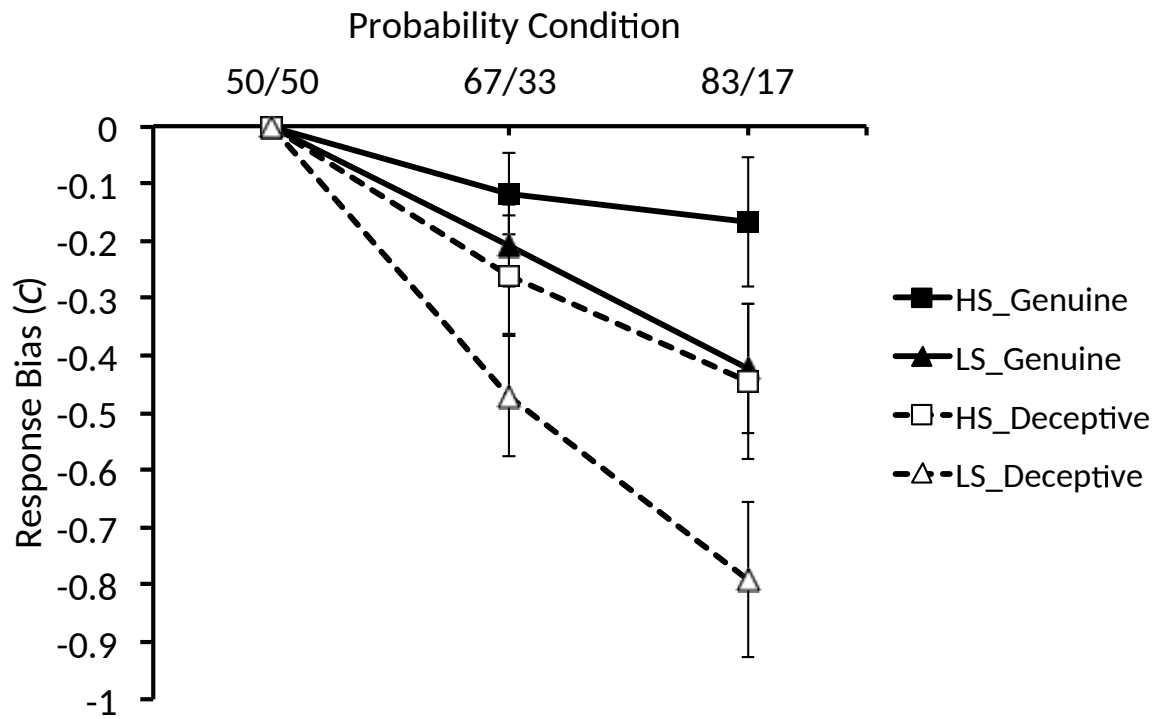




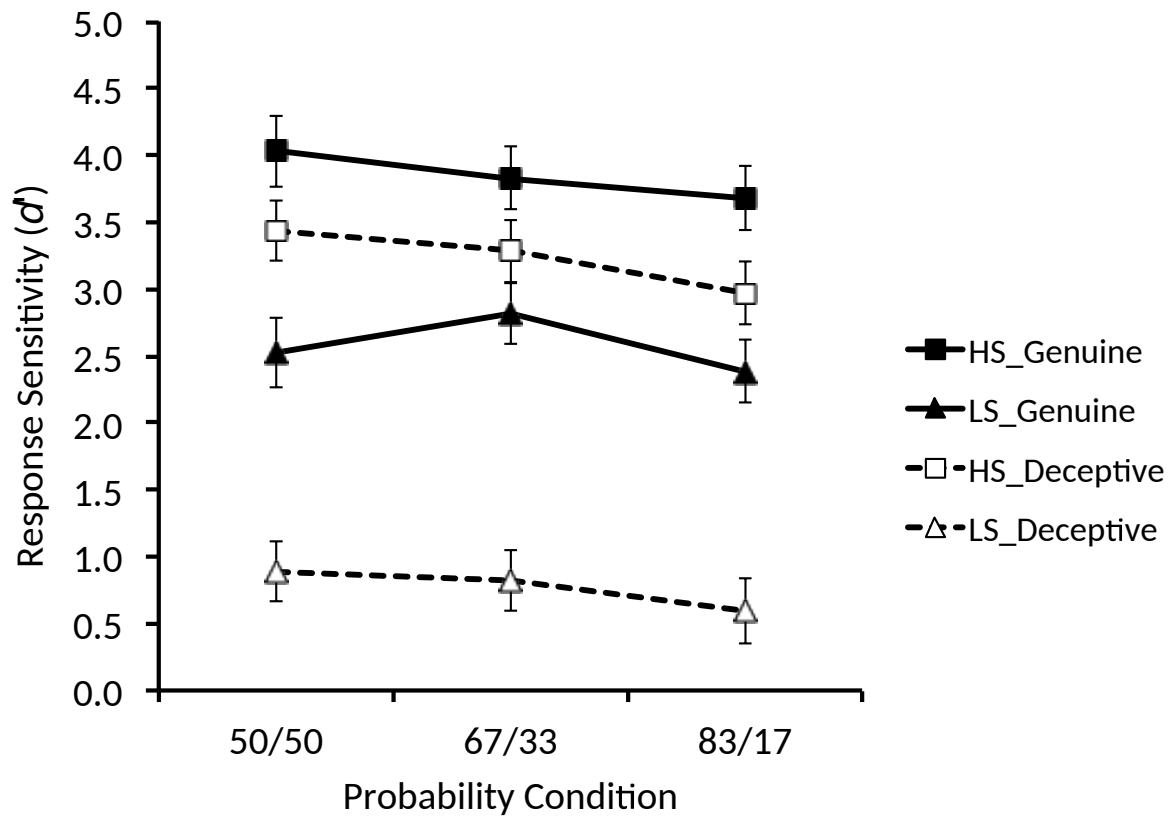
**Figure 2.** Sensitivity in discriminating between genuine and deceptive actions for the high-skilled and low-skilled group for each outcome probability value.



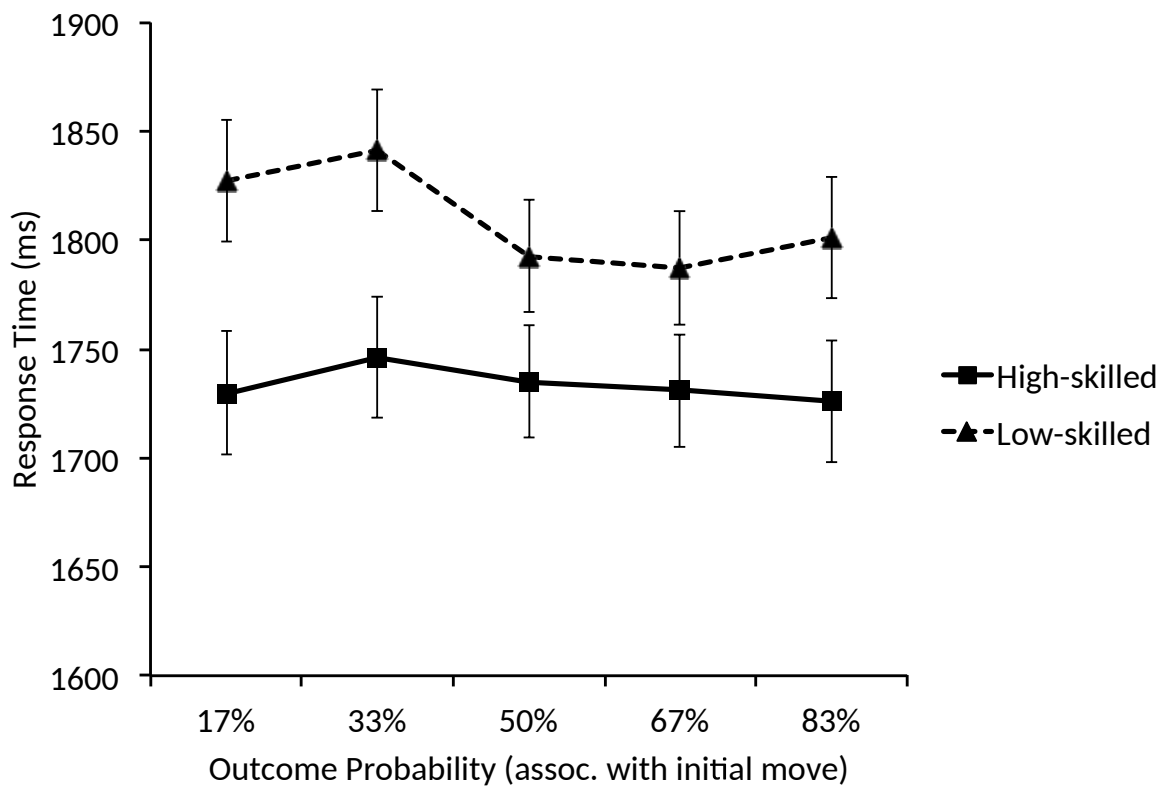
**Figure 3.** Bias toward judging actions to be genuine (–ve values) or deceptive (+ve values) for each outcome probability value.



**Figure 4.** Bias toward responding congruent with the higher-probability value (–ve bias) on genuine and deceptive trials in each outcome probability condition.



**Figure 5.** Sensitivity in discriminating between left and right action outcomes in the high-skilled (HS) and low-skilled (LS) group when viewing genuine and deceptive trials in each outcome probability pairing.



**Figure 6.** Mean response times for the high-skilled and low-skilled groups for each outcome probability.