## Measurement of time-dependent $C P$ asymmetries in $B^{0} \rightarrow D^{(*) \pm} \pi^{\mp}$ decays and constraints on $\sin (2 \beta+\gamma)$

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[^0]We present a measurement of $C P$-violating asymmetries in fully reconstructed $B^{0} \rightarrow D^{(*) \pm} \pi^{\mp}$ decays in approximately 88 million $\Upsilon(4 S) \rightarrow B \bar{B}$ decays collected with the BABAR detector at the PEP-II asymmetric-energy $B$ factory at SLAC. From a time-dependent maximum likelihood fit we obtain for the $C P$-violating parameters: $a=-0.022 \pm 0.038$ (stat.) $\pm 0.020$ (syst.), $a^{*}=$ $-0.068 \pm 0.038$ (stat.) $\pm 0.020$ (syst.), $c_{\text {lep }}=+0.025 \pm 0.068$ (stat.) $\pm 0.033$ (syst.), and $c_{\text {lep }}^{*}=+0.031 \pm$ 0.070 (stat.) $\pm 0.033$ (syst.). Using other measurements and theoretical assumptions we interpret

> the results in terms of the angles of the Cabibbo-Kobayashi-Maskawa unitarity triangle, and find $|\sin (2 \beta+\gamma)|>0.69$ at $68 \%$ confidence level. We exclude the hypothesis of no $C P$ violation ( $\sin (2 \beta+$ $\gamma)=0)$ at $83 \%$ confidence level.

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In the Standard Model, $C P$ violation in the weak interactions between quarks manifests itself as a non-zero area of the Cabibbo-Kobayashi-Maskawa (CKM) unitarity triangle 1]. While it is sufficient to measure one of its angles $\alpha, \beta$, or $\gamma$ to be different from 0 or $180^{\circ}$ to demonstrate the existence of $C P$ violation, the unitarity triangle needs to be overconstrained with different measurements to test the CKM mechanism. Measurements of $\beta$ free from theoretical uncertainties exist [2, 3], but there are no such measurements of $\alpha$ and $\gamma$. This letter reports the measurement of $C P$-violating asymmetries in $B^{0} \rightarrow D^{(*) \pm} \pi^{\mp}$ decays [4] in $\Upsilon(4 S) \rightarrow B \bar{B}$ decays and its interpretation in terms of constraints on $|\sin (2 \beta+\gamma)|$ [5, 6].

The time evolution of $B^{0} \rightarrow D^{(*) \pm} \pi^{\mp}$ decays is sensitive to $\gamma$ because of the interference between the CKMfavored decay $\bar{B}^{0} \rightarrow D^{(*)+} \pi^{-}$, whose amplitude is proportional to the CKM matrix elements $V_{c b} V_{u d}^{*}$, and the doubly-CKM-suppressed decay $B^{0} \rightarrow D^{(*)+} \pi^{-}$, whose amplitude is proportional to $V_{c d} V_{u b}^{*}$. The relative weak phase between the two amplitudes is $\gamma$, which, when combined with $B^{0} \bar{B}^{0}$ mixing, yields a weak phase difference of $2 \beta+\gamma$ between the interfering amplitudes.

The decay rate distribution for $B^{0} \rightarrow D^{ \pm} \pi^{\mp}$ decays is

$$
\begin{align*}
f^{ \pm}(\eta, \Delta t)= & \frac{e^{-|\Delta t| / \tau}}{4 \tau} \times  \tag{1}\\
& {\left[1 \mp S_{\zeta} \sin \left(\Delta m_{d} \Delta t\right) \mp \eta C \cos \left(\Delta m_{d} \Delta t\right)\right] }
\end{align*}
$$

where $\tau$ is the $B^{0}$ lifetime, neglecting the decay width difference, $\Delta m_{d}$ is the $B^{0} \bar{B}^{0}$ mixing frequency, and $\Delta t=t_{\mathrm{rec}}-t_{\mathrm{tag}}$ is the time of the $B^{0} \rightarrow D^{ \pm} \pi^{\mp}$ decay $\left(B_{\mathrm{rec}}\right)$ relative to the decay of the other $B\left(B_{\mathrm{tag}}\right)$. In this equation the upper (lower) sign refers to the flavor of $B_{\mathrm{tag}}$ as $B^{0}\left(\bar{B}^{0}\right)$, while $\eta=+1(-1)$ and $\zeta=+(-)$ for the final state $D^{-} \pi^{+}\left(D^{+} \pi^{-}\right)$. In the Standard Model, the $S$ and $C$ parameters can be expressed as

$$
\begin{equation*}
S_{ \pm}=-\frac{2 \operatorname{Im}\left(\lambda_{ \pm}\right)}{1+\left|\lambda_{ \pm}\right|^{2}}, \quad \text { and } \quad C=\frac{1-r^{2}}{1+r^{2}} \tag{2}
\end{equation*}
$$

where $\lambda_{ \pm}=r^{ \pm 1} e^{-i(2 \beta+\gamma \mp \delta)}$. Here $\delta$ is the relative strong phase and $r$ is the magnitude of the ratio of the suppressed and the favored amplitudes. The same equations apply for $B^{0} \rightarrow D^{* \pm} \pi^{\mp}$ decays, with $r$ and $\delta$ replaced by the parameters $r^{*}$ and $\delta^{*}$, respectively 7].

The analysis strategy is similar to that of the timedependent mixing measurement performed at $B A B A R$ [8]. To identify the flavor of $B_{\text {tag }}$, each event is assigned by a neural network to one of four hierarchical, mutually exclusive tagging categories: one lepton and two kaon categories based on the charges of identified leptons and
kaons, and a fourth category for remaining events. The effective tagging efficiency is $(28.1 \pm 0.7) \%$ [2]. The time difference $\Delta t$ is calculated from the separation along the beam collision axis, $\Delta z$, between the $B_{\mathrm{rec}}$ and $B_{\text {tag }}$ decay vertices. We determine the $B_{\text {rec }}$ vertex from its charged tracks. The $B_{\text {tag }}$ decay vertex is obtained by fitting tracks that do not belong to $B_{\mathrm{rec}}$, imposing constraints from the $B_{\text {rec }}$ momentum and the beam-spot location. The $\Delta t$ resolution is approximately 1.1 ps .

The expected $C P$ asymmetry in these decays is small $\left(r^{(*)} \approx\left|V_{u b}^{*} V_{c d} / V_{u d}^{*} V_{c b}\right| \approx 0.02\right)$, and therefore this measurement is sensitive to the interference between the $b \rightarrow u$ and $b \rightarrow c$ amplitudes in the decay of $B_{\mathrm{tag}}$. To account for this effect we use a parametrization different from Eq. 2] which is described in Ref. [9] and summarized here. For each tagging category $(i)$ the interference is parametrized in terms of the effective parameters $r_{i}^{\prime}$ and $\delta_{i}^{\prime}$. Neglecting terms of order $r^{(*) 2}$ and $r_{i}^{\prime 2}$, for each tagging category the $\Delta t$ distribution becomes

$$
\begin{align*}
f_{i}^{ \pm(*)}(\eta, \Delta t)= & \frac{e^{-|\Delta t| / \tau}}{4 \tau} \times\left[1 \mp\left(a^{(*)} \mp \eta b_{i}-\eta c_{i}^{(*)}\right)\right. \\
& \left.\sin \left(\Delta m_{d} \Delta t\right) \mp \eta \cos \left(\Delta m_{d} \Delta t\right)\right] \tag{3}
\end{align*}
$$

where, in the Standard Model,

$$
\begin{align*}
a^{(*)} & =2 r^{(*)} \sin (2 \beta+\gamma) \cos \delta^{(*)} \\
b_{i} & =2 r_{i}^{\prime} \sin (2 \beta+\gamma) \cos \delta_{i}^{\prime} \\
c_{i}^{(*)} & =2 \cos (2 \beta+\gamma)\left(r^{(*)} \sin \delta^{(*)}-r_{i}^{\prime} \sin \delta_{i}^{\prime}\right) \tag{4}
\end{align*}
$$

Semileptonic $B$ decays do not have a doubly-CKMsuppressed amplitude contribution, and hence $r_{\text {lep }}^{\prime}=0$. Given that we have two $B$ decay modes and four tagging categories, we use two $a$ parameters (one for each final state), three $b$ parameters (one for each non-lepton tagging category), and eight $c$ parameters (one for each combination of tagging category and final state). Results are quoted only for the four parameters $a^{(*)}$ and $c_{\text {lep }}^{(*)}$, which are independent of the unknowns $r_{i}^{\prime}$ and $\delta_{i}^{\prime}$. The other parameters are allowed to float in the fit, but, as they depend on $r_{i}^{\prime}$ and $\delta_{i}^{\prime}$, they do not contribute to the interpretation of the result in terms of $\sin (2 \beta+\gamma)$.

This measurement is based on 88 million $\Upsilon(4 S) \rightarrow$ $B \bar{B}$ decays, corresponding to an integrated luminosity of $82 \mathrm{fb}^{-1}$, collected with the BABAR detector [10] at the PEP-II asymmetric-energy $B$ factory at SLAC. We use a Monte Carlo simulation of the BABAR detector based on GEANT4 11] to validate the analysis procedure and to estimate some of the backgrounds.

The event selection and the reconstruction of $B^{0} \rightarrow$ $D^{(*) \pm} \pi^{\mp}$ candidates are detailed in Ref. 8]. Signal


FIG. 1: Distributions of $m_{\mathrm{ES}}$ in the $\Delta E$ signal region for events with tagging information in the $B^{0} \rightarrow D^{ \pm} \pi^{\mp}$ (left plot) and the $B^{0} \rightarrow D^{* \pm} \pi^{\mp}$ sample (right plot).
and background are discriminated by two kinematic variables: the beam-energy substituted mass, $m_{\mathrm{ES}} \equiv$ $\sqrt{(\sqrt{s} / 2)^{2}-p_{B}^{* 2}}$, and the difference between the $B$ candidate's measured energy and the beam energy, $\Delta E \equiv$ $E_{B}^{*}-(\sqrt{s} / 2)$, where $E_{B}^{*}\left(p_{B}^{*}\right)$ is the energy (momentum) of the $B$ candidate in the $e^{+} e^{-}$center-of-mass frame, and $\sqrt{s}$ is the total center-of-mass energy. The signal region is defined as $|\Delta E|<3 \sigma$, where the resolution $\sigma$ is mode-dependent and approximately 20 MeV , as determined from data. Figure 1 shows the $m_{\mathrm{ES}}$ distribution for candidates in the $\Delta E$ signal region. The $m_{\mathrm{ES}}$ distribution is fit with the sum of a threshold function 12], which accounts for the background from random combinations of tracks, and a Gaussian distribution with a fitted width of about $2.5 \mathrm{MeV} / c^{2}$ describing the signal. After tagging, the Gaussian yield is $5207 \pm 87$ and $4746 \pm 78$ events for the $B^{0} \rightarrow D^{ \pm} \pi^{\mp}$ and $B^{0} \rightarrow D^{* \pm} \pi^{\mp}$ sample respectively, with corresponding purities of $(84.9 \pm 0.5) \%$ and ( $94.4 \pm 0.4) \%$ in a $\pm 3 \sigma$ region around the nominal $B$ mass. Backgrounds from $B^{0}$ decays that peak in the $m_{\mathrm{ES}}$ signal region were estimated with Monte Carlo simulation to constitute $(0.21 \pm 0.06) \%$ and $(0.13 \pm 0.05) \%$ of the $B^{0} \rightarrow D^{ \pm} \pi^{\mp}$ and $B^{0} \rightarrow D^{* \pm} \pi^{\mp}$ yields, respectively. For backgrounds from $B^{+}$decays, the corresponding figures are $(0.93 \pm 0.23) \%$ and $(0.93 \pm 0.10) \%$.

An unbinned maximum-likelihood fit is performed on the selected $B$ candidates using the $\Delta t$ distribution in Eq. 3 convolved with a resolution function composed of three Gaussian distributions. Incorrect tagging dilutes the parameters $a^{(*)}, c_{i}^{(*)}$, and the coefficient of $\cos \left(\Delta m_{d} \Delta t\right)$ by a factor $D_{i}=1-2 w_{i}$ [2, 19], where $w_{i}$ is the mistag fraction. The resolution function and the parameters associated with flavor tagging are determined from the data and are consistent with previous $B A B A R$ analyses [2]. The combinatorial background is parametrized as the sum of a component with zero lifetime and one with an effective lifetime fixed to the value obtained from simulation. The fraction of each component and the $\Delta t$ resolution parameters are left free in the fit to the data. The background coming from $B^{ \pm}$ mesons is modeled with an exponential decay with the $B^{ \pm}$lifetime, and its size is fixed to the value predicted
by simulation. The background from $B^{0}$ mesons is neglected in the nominal fit, but is considered in evaluating the systematic uncertainties.

The results from the fit to the data are

$$
\begin{align*}
a & =-0.022 \pm 0.038 \text { (stat.) } \pm 0.020 \text { (syst.) } \\
a^{*} & =-0.068 \pm 0.038 \text { (stat.) } \pm 0.020 \text { (syst.) } \\
c_{\text {lep }} & =+0.025 \pm 0.068 \text { (stat.) } \pm 0.033 \text { (syst.) } \\
c_{\text {lep }}^{*} & =+0.031 \pm 0.070 \text { (stat.) } \pm 0.033 \text { (syst.) } \tag{5}
\end{align*}
$$

All other fitted $b$ and $c$ parameters are consistent with zero. Figure 2 shows the fitted $\Delta t$ distributions for events from the lepton tagging category, which has the lowest level of background and mistag probability.

The systematic uncertainties on the parameters in Eq. 5 has been calculated in a manner similar to that used in Ref. [8]. A small bias in the $\Delta t$ measurement could result in a bias on the $c$ parameters in Eq. 3 For instance, a realistic $\Delta t$ bias of 0.024 ps results in a shift in $c_{\text {lep }}^{*}$ of 0.002 . We are immune from this effect because we fit for tagging category dependent biases in the resolution function directly on data. Nonetheless, the impact of a possible mismeasurement of $\Delta t$ has been estimated by varying the assumptions on the resolution function, the position of the beam-spot, the absolute $z$ scale, the internal alignment of the vertex detector, and quality criteria on the reconstructed vertex. The corresponding error on $a^{(*)}$ is $\sigma_{a}=0.015$, while that on $c^{(*)}$ is $\sigma_{c}=0.026$. The systematic uncertainties on the fit technique ( $\sigma_{a}=0.013$, $\left.\sigma_{c}=0.020\right)$ include the upper limit on the fit bias estimated from samples of fully simulated events, the uncertainty on the $B^{0}$ lifetime and $\Delta m_{d}$ [13], and the impact of neglecting higher order terms in $r^{(*)}$ or $r_{i}^{\prime}$ in Eq. 3 As a cross-check, we performed the same fits on samples of $18233 B^{-} \rightarrow D^{(*) 0} \pi^{-}$and $1740 \bar{B}^{0} \rightarrow J / \psi K^{* 0}$ candidates, where we find no significant $C P$ asymmetries, as expected. The systematic uncertainties in tagging ( $\sigma_{a}=0.004, \sigma_{c}=0.003$ ) are estimated allowing for different tagging efficiencies between $B^{0}$ and $\bar{B}^{0}$ and for different $\Delta t$ resolutions for correctly and incorrectly tagged events. We also account for uncertainties on the background ( $\sigma_{a}=0.001, \sigma_{c}=0.003$ ) by varying the effective lifetimes, dilutions, $m_{\mathrm{ES}}$ shape parameters and signal fractions, and background $C P$ asymmetry up to five times the expected $C P$ asymmetry for signal.

The results can be interpreted in terms of $\sin (2 \beta+\gamma)$ (Eq. (4) if the decay amplitude ratios $r^{(*)}$, expected to be $\left|V_{u b}^{*} V_{c d} / V_{u d}^{*} V_{c b}\right| \approx 0.02$, are known. Such small amplitude ratios cannot be determined from $B^{0} \rightarrow D^{(*) \pm} \pi^{\mp}$ events directly, because the current data sample is too small. We estimate $r^{(*)}$ using the $S U(3)$ symmetry relation $r^{(*)}=\tan \theta_{C} \sqrt{\frac{\mathcal{B}\left(B^{0} \rightarrow D_{s}^{(*)+} \pi^{-}\right)}{\mathcal{B}\left(B^{0} \rightarrow D^{(*)-} \pi^{+}\right)}} \frac{f_{D(*)}}{f_{D_{s}(*)}^{(*)}}$ [5]. From the measurements of the Cabibbo angle $\tan \theta_{C}=0.2250 \pm$ 0.0027 13], the branching fractions $\mathcal{B}\left(B^{0} \rightarrow D^{-} \pi^{+}\right)$ $=(0.30 \pm 0.04) \%$ 13], $\mathcal{B}\left(B^{0} \rightarrow D^{*-} \pi^{+}\right)=(0.276 \pm$


FIG. 2: Distributions of $\Delta t$ for the $B^{0} \rightarrow D^{ \pm} \pi^{\mp}(\mathrm{a}-\mathrm{d})$ and $B^{0} \rightarrow D^{* \pm} \pi^{\mp}$ (e-h) candidates tagged with leptons, split by $B$ tagging flavor and reconstructed final state. The lines are fit projections and hatched regions represent background.
$0.021) \%$ [13], $\mathcal{B}\left(B^{0} \rightarrow D_{s}^{+} \pi^{-}\right)=\left(2.7_{-0.6}^{+0.7} \pm 0.8\right) \times 10^{-5}$ 14], $\mathcal{B}\left(B^{0} \rightarrow D_{s}^{*+} \pi^{-}\right)=\left(1.9_{-1.3}^{+1.2} \pm 0.5\right) \times 10^{-5}$ [14] , and from calculations of the decay constant ratios $f_{D_{s}} / f_{D}$ $=1.11 \pm 0.01$ and $f_{D_{s}^{*}} / f_{D^{*}}=1.10 \pm 0.02$ 15] we obtain

$$
\begin{equation*}
r=0.019 \pm 0.004, \quad r^{*}=0.017_{-0.007}^{+0.005} \tag{6}
\end{equation*}
$$

To obtain $\sin (2 \beta+\gamma)$, we minimize the $\chi^{2}$

$$
\begin{equation*}
\chi^{2}\left(2 \beta+\gamma, \delta^{(*)}, r^{(*)}\right)=\sum_{i}\left(\frac{\tilde{x}_{i}-x_{i}}{\sigma_{i}}\right)^{2}+\Delta\left(r^{(*)}\right) \tag{7}
\end{equation*}
$$

where $x_{i}=a, a^{*}, c_{\text {lep }}, c_{\text {lep }}^{*}$ are functions of the physics parameters (Eq.4), and $\tilde{x}_{i}$ are the corresponding measured values. $\Delta\left(r^{(*)}\right)$ is a continuous function that is set equal to 0 within $30 \%$ of the estimated $r^{(*)}$ (Eq. 6), and is an offset quadratic outside this range, with the errors in Eq.6] The additional $30 \%$ error attributed on $r^{(*)}$ is due to the unknown theoretical uncertainty on the validity of the $S U(3)$ symmetry assumption and to neglecting $W$ exchange contributions to $A\left(B^{0} \rightarrow D^{(*)+} \pi^{-}\right)$. This error estimate is consistent with the spread in $r^{(*)}$ obtained using a variety of theoretical models 16]. The $\sigma_{i}$ are the quadratic sums of the statistical and systematic uncertainties in Eq. [5] Correlations between the $\tilde{x}_{i}$, at most $28 \%$, have negligible influence on the results of this analysis. The simultaneous analysis of two $B$ decay modes allows one to extract $|\sin (2 \beta+\gamma)|$.

Figure 3 shows the minimum $\chi^{2}$ for each value of $|\sin (2 \beta+\gamma)|$. The absolute minimum occurs for $\mid \sin (2 \beta+$ $\gamma) \mid=0.98$, where $\chi_{\min }^{2} /$ d.o.f. $=0.44 / 1$. The values of $r^{(*)}$ that minimize the $\chi^{2}$ are consistent with the input values within their statistical errors. Because of the large uncertainties on the fit parameters and their limited physical range, the $\chi^{2}$ curve is non-parabolic. Thus to obtain a probabilistic interpretation to the results, we consider, for each of many values of $\sin (2 \beta+\gamma)$, a large number of simulated experiments with the same characteristics as the data. We compute the consistency of the data


FIG. 3: Dependence of $\chi^{2}$ on $|\sin (2 \beta+\gamma)|$ (top) and of the frequentist confidence level of the agreement of the data with expectations as a function of the hypothesis on $|\sin (2 \beta+\gamma)|$ (bottom). The assumptions on $r$ and $r^{*}$ are contained in the definition of $\chi^{2}$ (Eq. 7). The dashed horizontal lines indicate the $68 \%$ and $83 \%$ confidence levels (defined in the text).
with a given value of $\sin (2 \beta+\gamma)$ by counting the fraction of simulated experiments for which $\chi^{2}(\sin (2 \beta+\gamma))-\chi_{\text {min }}^{2}$ is smaller than it is in the data. This fraction, the frequentist confidence level, is shown in the lower portion of Fig. 3 from which we read that $|\sin (2 \beta+\gamma)|>0.69$ at $68 \%$ C.L. We exclude the hypothesis of no $C P$ violation $(\sin (2 \beta+\gamma)=0)$ at $83 \%$ confidence level. In order to study the impact of the assumed theoretical error on $r^{(*)}$, we doubled it to $60 \%$ and we found that the lower limit on $|\sin (2 \beta+\gamma)|$ at $68 \%$ C.L. drops from 0.69 to 0.60 .

In conclusion, we studied the time-dependent $C P$-violating asymmetries in fully reconstructed $B^{0} \rightarrow D^{(*) \pm} \pi^{\mp}$ decays, and measured the $C P$-violating parameters listed in Eq. 5. With some theoretical assumptions, we interpret the result in terms of $\sin (2 \beta+\gamma)$ and we find that $|\sin (2 \beta+\gamma)|>0.69$ at $68 \%$ C.L. and that $\sin (2 \beta+\gamma)=0$ is excluded at $83 \%$ C.L.

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